SYSTEMATIC REVIEW



Bite force of children and adolescents: a systematic review and meta-analysis

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1. Introduction

Abstract

This systematic review aimed to assess bite force measurements in children and adolescents and to study the various devices that measure Maximum Voluntary Bite Force (MVBF). This systematic review included observational studies and experimental studies in children and adolescents (upto 19 years of age) which evaluated MVBF using a bite force measuring device. Studies on participants with systemic conditions were excluded. Databases such as PubMed, Embase, LILACS, and the Cochrane library were searched until September 2022, for which screening and quality assessment were performed. Newcastle-Ottawa, modified Newcastle-Ottawa and ROBINS-I tools were used to assess the Risk-of-bias. All observational studies reporting overall bite force values of participants were included for meta-analyses. A total of 8864 participants (3491 males and 3623 females) were included from 61 studies. Meta-analyses were conducted to evaluate mean average bite force value for each included dentition using R software v2.4-0. Estimation was done to derive an average BF value for variables such as age (dentition), gender, side, site, device and ethnicity. MVBF values were reported as mean average in the form of MLN with 95% CI (Confidence Interval). Using a randomeffects model, 29 forest plots were generated. I² values varied between 90% and 100%. Bite force ranged from 246.22 N (220.47; 274.98) to 311.72 N (255.99; 379.59) and 489.35 N (399.86; 598.87) in primary, mixed, and permanent dentitions, respectively. Six different sites for recording bite force and 11 different types of devices were reported with portable occlusal bite force gauge being the most common device. Outcomes of this review provide useful baseline reference values of bite force for clinicians and researchers.

Keywords

Bite force; Children; Device; Dentition; Occlusion

Bite force (BF) can be defined as "the capacity of the mandibular elevation muscles to perform a maximum force of lower teeth against the upper teeth, under favorable conditions" [1, 2]. The measurement of these balanced forces such as Maximum Bite Force (MBF) or Maximum Voluntary Bite Force (MVBF) and their dispersion could be an index for assessment of the level of normality or deviation from the normal in oral health. In any compromised dentition, the assessment of occlusion and the forces applied within the stomatognathic system when in occlusion may aid in the quantification of the clinical complications in an individual. BF measurements have been studied in various subjects with a wide range of devices for either the diagnosis or assessment of multiple conditions such as dental decay [3, 4], temporomandibular joint or muscle disorders [5, 6], malocclusion [7, 8], influence of Body Mass Index (BMI) [9], and early loss of teeth, which can disturb normal occlusion.

BF values can be obtained directly or indirectly from an individual [10, 11]. Among the many parameters used for assessing BF by clinical measures are Masticatory Efficiency or Performance [12], Electromyographic (EMG) activity [13], and measurement of MVBF [14, 15]. However, unlike MVBF, masticatory performance and EMG are not quantifiable as numerical data. Additionally, reports in the literature have associated MVBF and its impact on quality of life directly through multiple studies in children [16–18]. A wide range of MVBF values has been recorded, and numerous devices for recording them have been reported in the literature thus far. A recent review on BF measuring devices [19] listed nine different types of equipment commonly used in gauging MVBF in children, adolescents, and adults.

Uncertainty prevails regarding the calibrated values of MVBF, measured in children and adolescents due to multiple devices, varying sites and units used to evaluate BF. Also, the absence of a baseline value in primary, mixed and permanent dentitions necessitates a strong need for quantitative evaluation



FIGURE 1. Flow diagram for study selection according to the PRISMA (Preferred Reporting Items for Systematic **Reviews and Meta-Analyses**) guidelines. Total of 57 studies that satisfied the eligibility criteria were included in this systematic review.

of the average BF value. Values that may serve as standard baseline measurements to detect various dental-related abnormalities concerning age/dentition, gender, side, site, device and ethnicity are required. Therefore, this systematic review is primarily intended to compile BF measurements in children and adolescents and investigate various devices that measure MVBF. Additionally, this systematic review aimed to document changes in BF after various methods of oral rehabilitation, significant to diagnose any clinical deviations from the baseline value.

2. Methods

This systematic review was performed to review BF measurements in children and adolescents and reported according to PRISMA (Preferred Reporting Items for Systematic Reviews and Meta-Analyses) guidelines [20]. The study protocol was registered in PROSPERO (The International Prospective Register of Systematic Reviews) no: CRD42020150464.

2.1 Selection criteria

This systematic review included observational studies (crosssectional, case-control, and cohort studies) and experimental studies evaluating MVBF using a BF measuring device. Children and adolescents (according to WHO (World Health Organization criteria)) [21] comprised the study population. Study designs such as reviews, protocols for trials, and conference abstracts were excluded. Any studies that measured MVBF in adults (>19 years), children with special health care needs, with cleft lip/palate, and medically compromised patients were excluded. Records published in languages other than English and unpublished records were excluded.

2.2 Search strategy

The search aimed to find published data with no limit on the date of publication up to September 2022 from PubMed, Embase (Excerpta Medica Database), LILACS (Latin American & Caribbean Health Sciences Literature), and Cochrane library. Two authors (PJ and GFP) independently identified potential references through hand searches of journals (Journal of Dental Research, Paediatric Dentistry, Journal of Dentistry for Children, International Journal of Pediatric Dentistry, International Journal of Orthodontics, Journal of Clinical Orthodontics, Journal of Prosthodontics, and Journal of Prosthetic Dentistry), cross-references, and grey literature (NTIS (National Technical Information Services), Ovid, PsycEX-TRA, and Shodhganga) for relevant articles. Search terms were a combination of MeSH and free text such as "bite force", "masticatory efficiency", "occlusal force distribution", "child", "children", "adolescents", "device", "method", and "measurement".

2.3 Data collection

Two reviewers (PJ and GFP) independently performed an initial title and abstract screening to select the articles according to the eligibility criteria. The consensus was sought from a third investigator (MS Muthu) in cases of disagreement about the eligibility of a study to be retained. All studies excluded at this stage are documented with reasons for exclusion (not satisfying inclusion criteria, no study-related outcome, insufficient data despite attempts to contact the author, and not addressing the review question). Two authors (PJ and GFP) independently extracted data from each included study. The customized spreadsheets included the details on author/s, year of publication, demographic information of the participants (age, gender, ethnicity), dentition, index test (masticatory efficiency/bite force/occlusal force distribution) used, instrument/device used, and unit of measurement along with side and site of bite registration. Further information was sought from the authors of the studies wherever necessary.

2.4 Assessment of risk-of-bias

The risk-of-bias in the included studies was assessed by PJ and GFP independently. Assessing the risk-of-bias among included cohort and case-control studies was done using the Newcastle-Ottawa scale (NOS) [22] and the modified Newcastle-Ottawa scale (M-NOS) [23] for cross-sectional studies. The risk-of-bias was categorized as high (\geq 7 stars), moderate (5–6 stars), and low (<5 stars) based on the NOS and M-NOS scales. Additionally, the ROBINS-I (Risk of Bias in Non-randomised Studies-of Interventions) tool [24] judged the quality among included non-randomized clinical trials as critical, serious, or low. The third author (MSM) resolved any differences in methodological quality assessments. We sought further information from the authors of the studies wherever data were inadequate.

2.5 Statistical analysis

The data extracted were entered into customized spreadsheets. Studies that reported average MVBF values and standard deviation were categorized into primary (3-6 years), mixed (6-12 years), permanent (12-19) and overlapping dentitions (all children, whatever their age or dentition development stage) and pooled for meta-analysis utilizing the R Software metafor package v2.4-0 (Wolfgang Viechtbauer, Boston, MA, USA) [25]. Forest plots were generated for parameters of concern, to represent the estimated effect of mean BF across dentitions. Subgroup analysis was done for BF values based on gender, side, site, measuring devices and ethnicity. Random effects model measured the average differences in the investigated outcome. Statistical heterogeneity was estimated by the chisquare test (p < 0.05) and the I-square index (I²). The classification of I² values were, over 50% as high, 25–50% moderate, and <25% as low [26].

3. Results

Subsequent to a literature search, 614 unique records were identified for this systematic review. Following the elimination of 40 duplicates, 574 studies were appraised for abstracts, of which 366 studies (studies that did not match the inclusion criteria such as, participants beyond 19 years of age, medically compromised and special children, TMJ and muscle disorders, animal studies and finite element analysis) were excluded, resulting in 208 articles for full-text reading. Two publications from a supplement journal could not be retrieved, leading to a total of 206 records. After scrutiny, 145 articles from the 206 records were excluded for the aforementioned reasons—not satisfying inclusion criteria (n = 34), no study-related outcome (n = 12), insufficient data despite attempts to contact the author (n = 82), and not addressing the review question (n = 17). Therefore, 61 studies that satisfied the eligibility criteria were included in this systematic review (Fig. 1).

3.1 Study characteristics (Supplementary Table 1)

Among the 61 included studies, publications dating from 1986 to 2022 were included. Study designs were divided into cohort (n = 2), case-control (n = 1), cross-sectional (n = 48), and clinical trials (n = 10). Of the data points extracted, 22 unique world regions were identified (Brazil, Croatia, Colombia, Denmark, Ireland, Greece, Japan, Jordan, Kingdom of Saudi Arabia, Iran, Iraq, India, Netherlands, South Korea, Switzerland, Istanbul, Taiwan, Thailand, UK, Poland, Sweden, and the US), with the largest contributions from Brazil, Japan, and Jordan. Ages ranged from 3 to 19 years, with a total of 8864 included participants. Of the 61 included studies, only 26 studies had BF information based on gender, both males (n = 3491) and females (n = 3623), while the one study [27] had information on males only (n = 34). The remaining 34 studies included 2124 participants with no mention of gender ratio.

3.2 Quality of reviewed studies

Table 1 and Fig. 2 represents the risk-of-bias for the included observational studies and clinical trials respectively. The majority had a low risk-of-bias (n = 31), with moderate risk in 16 publications and four reports with high risk-of-bias according to the NOS and M-NOS scales. Among the ten clinical trials, based on the ROBINS-I tool, seven studies had critical risk and three had serious risk. The studies scored with a high risk-of-bias had inadequate information in multiple domains, such as sample size and ascertainment of exposure.

3.3 Data categorization

Quantitatively assessment of 42 eligible reports included studies measuring BF in either one of the dentitions exclusively or in combinations of two or more dentitions. In total, 13, 18, and 11 studies reported BF values distinctively for the primary, mixed, and permanent dentitions, respectively. **Supplementary Table 1** tabulates the MVBF measurements assessed by various devices and reports the maximum BF values (mean average) in the form of MLN (log-transformed mean) with 95% CI. I² values varied between 90% and 100%, indicating substantial heterogeneity, resulting in 29 forest plots (Figs. 3,4,5) through random effects model.

3.3.1 Primary dentition (ages 3 to 6 years)

Of 13 studies included for quantitative synthesis, 12 primary dentition studies computed the average BF as 246.22 (220.47; 274.98) Newtons (N) from 2,155 children. For gender assessment (Fig. 4A–B), the molar BF reported in 751 boys—222.30 (187.90; 263.01) N, showed a slight increase in BF values compared to 759 girls (202.43 (182.78; 224.19) N). The average BF values of the right and the left sides were 169.79 (129.31; 222.93) N and 163.88 (125.22; 214.48) N,

(A) Ris	k of bias for cross sectional st	udies						
S. No	Author Name		Selection		Comparability	Outc	ome	No. of Stars
	(Year)		(Max 5 stars	s)	(Max 2 stars)	(Max 3	3 stars)	Risk of Bias
1	Zwir [6]	*	*	**	**	**	*	9
	(2018)							Low
2	Takeshima [83]	*	*	*	**	*	*	7
	(2019)							Low
3	Abu Alhaija [30]	*	*	*	**	*	*	7
	(2018)							Low
4	Kiriishi [27]	*	*	_	-	**	*	5
	(2018)							Moderate
5	Heydari [31]	*	*	-	**	*	*	6
	(2018)							Moderate
6	Jeong [70]	*	*	*	**	*	*	7
	(2019)							Low
7	Pereira [75]	*	*	-	*	*	*	5
	(2018)							Moderate
8	Diaz Serrano [65]	*	*	**	*	**	*	8
-	(2017)							Low
9	Hama $\begin{bmatrix} 68 \end{bmatrix}$	*	_	**	*	*	*	6
2	(2017)							Moderate
10	(2017) Kava [54]	*	*	*	**	**	*	8
10	(2017)							Low
11	(2017) Awawdeh [55]	*	*	*	**	*	*	7
11	(2017)							Low
12	(2017) Marquezin [73]	*	*	**	_	**	*	7
12	(2016)				-			Low
12	(2010) A raujo [52]	*	*	**	**	**	*	LOW 0
15	(2016)							y Low
14	(2010)	*	*	**	**	**	*	LOW
14	SZYINANSKA [82]		·				·	9
15	(2013)	*	*	**	**	*	*	LOW
15	Sun [9]		-					ð
16	(2016)	*	*	*	*	*	*	Low
16	Al Quassar [61]	Υ.	*	Ť	*	*	т	6
17	(2017)	.				ale	ste	Moderate
17	Sato [79]	*	-	-	-	*	*	3
10	(2011)			44 - 44		-1-	.t.	High
18	Varga [15]	-	-	**	*	*	*	5
	(2011)							Moderate
19	Mountain [29]	*	*	**	*	*	*	7
	(2010)							Low
20	Castelo [57]	*	-	**	**	*	*	7
	(2010a)							Low
21	Castelo [64]	*	*	**	*	-	*	6
	(2010b)							Moderate
22	Oueis [74]	*	*	**	*	*	*	7
	(2009)							Low

TABLE 1. Risk of bias assessment table for observational studies.

(A) Risk	of bias for cross sectional stu	idies						
S. No	Author Name		Selection		Comparability	Oute	come	No. of Stars
	(Year)		(Max 5 stars	5)	(Max 2 stars)	(Max)	3 stars)	Risk of Bias
22	Oueis [74]	*	*	**	*	*	*	7
	(2009)							Low
23	Thongudomporn [84]	*	*	**	**	*	*	8
	(2009)							Low
24	Usui [32]	*	*	*	**	*	*	7
	(2007)							Low
25	Pereira [76]	*	*	*	**	*	*	7
	(2007)							Low
26	Duarte Gaviao [51]	*	*	*	**	*	*	7
	(2006)							Low
27	Sakashita [78]	*	*	*	*	*	*	6
	(2006)							Moderate
28	Bonjardim [63]	*	*	*	**	-	*	6
	(2005)							Moderate
29	Sonnensan [81]	*	*	*	*	*	*	7
	(2005)							Low
30	Kamegai [41]	*	*	-	*	*	*	5
	(2005							Moderate
31	Maki [72]	*	*	-	*	*	*	5
	(2001)							Moderate
32	Kampe [14]	*	-	-	*	*	*	4
	(1987)							High
33	Fields [67]	*	-	-	*	*	*	4
	(1986)							High
34	Gaviao [76]	-	*	*	*	*	*	5
	(2007)							Moderate
35	Medhat [53]	*	*	*	**	*	*	7
	(2018)							Low
36	Marquezin [58]	*	*	**	**	*	*	8
	(2017)							Low
37	Haritha [69]	*	-	*	*	*	*	5
	(2012)							Moderate
38	Lemos [71]	*	-	*	**	*	*	6
	(2006)							Moderate
39	Karibe [28]	*	-	-	*	*	*	4
	(2003)							High

TABLE 1. Continued.

(A) Risl	x of bias for cross secti	onal stu	dies							
S. No	Author Name		Selection		Comparability	Outcom	e	No. of Stars		
	(Year)		(Max 5 stars)		(Max 2 stars)	(Max 3 sta	ars)	Risk of Bias		
40	Rentes [77]	*	-	*	*	*	*	5		
	(2002)							Moderate		
41	Sonnensen [56]	*	*	*	**	*	*	7		
	(2001)							Low		
42	Su [17]	*	*	**	**	-	*	7		
	(2009)							Low		
43	Alam [62]	*	*	*	**	*	*	7		
	(2020)							Low		
44	Sonnaville [80]	*	*	**	*	*	*	7		
	(2021)							Low		
45	Guo [86]	*	*	*	*	**	*	7		
	(2021)							Low		
46	Aishwaryaa [87]	*	*	-	**	*	*	6		
	(2021)							Moderate		
47	Prabahar [88]	*	*	*	**	**	*	8		
	(2021)							Low		
48	Gudipaneni [85]	*	*	**	*	**	*	8		
	(2020)							Low		
(B) Risk	c of bias for case contro	ol study								
S.No	Author Name		Selection		Comparability	Exposur	e	No. of Stars		
	(Year)		(Max 4 stars)		(Max 2 stars)	(Max 3 sta	ars)	Risk of Bias		
49	Ferraira [10]	*	_ *	*	**	* *	-	7		
	(2016)							Low		
(C) Risk	c of bias for cohort stud	dies								
S.No	Author Name		Selection		Comparability	Outcom	e	No. of Stars		
	(Year)		(Max 4 stars)		(Max 2 stars)	(Max 3 sta	ars)	Risk of Bias		
50	Roldan [8]	*		*	**	* *	*	7		

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*

*

Low 9

Low

*

TABLE 1. Continued.

51

*

(2016)

Ohira [12]

(2012)



FIGURE 2. Risk-of-bias assessment for clinical trials. Among the 10 clinical trials, based on the ROBINS-I tool, seven studies had critical risk and three had serious risk.

respectively [28-31].

A total of 7 different devices assessed the BF at the primary molar site. Quantitative analysis included two studies with devices comprising similar sites and units. From studies using the portable occlusal bite force (OBF) gauge (n = 4), BF averaged 195.33 (188.44; 202.47) N and from studies testing BF using the pressurized transducer (n = 3), the BF value estimated to 265.57 (233.94; 301.57) N (Fig. 5A).

Considering ethnicity, both Japan and Brazil reported an equal number of studies (n = 6) computing 204.58 (177.24; 236.13) N and 265.56 (233.94; 301.46) N, respectively (Fig. 5D).

3.3.2 Mixed dentition (ages 6 to 12 years)

In total, 18 authors recorded BF in the mixed dentition from 2,736 participants, which averaged 311.72 (255.99; 379.59) N. No clinically significant differences were observed in the molar BF of males (n = 763), 256.27 (156.91; 418.54) N, and females (n = 781), 241.45 (146.00; 399.28) N. The estimations on right and left molars averaged 160.42 (106.61; 241.39) N and 143.95 (94.92; 218.30) N, respectively.

BF values measured in kilograms (kg) (n = 208) and kilogram-force (kgf) (n = 140) showed that males had increased masticatory pressure compared with the females of the same group (Maki *et al.* [72] 2001). MVBF of males (n = 102) computed to 24.59 (20.29; 29.79) kg, and that



Study	Total	Mean	SD		Mean	Ме	an	9	5%–CI	Weight (common)	Weight (random)
Castelo1, 2010	13	352.81	23.6700	i	Ŧ	352	81	[340.17; 3	865.91]	6.6%	5.3%
Castelo2, 2010	14	333.02	44.1800		+-	333	02	[310.66; 3	856.99]	1.8%	5.3%
Kamegai1, 2005	486	296.30	128.1000			296	30	[285.13; 3	807.91]	6.0%	5.3%
Serra, 2007	22	302.00	61.0000	1 -	-	302	00	[277.56; 3	828.60]	1.2%	5.2%
Ravi kumar, 2020	120	320.93	54.2300		<u>in</u>	320	93	[311.37; 3	30.78]	9.6%	5.3%
Duarte Gaviao, 2006	15	441.63	44.3000			441	63	[419.77; 4	64.63]	3.4%	5.3%
Diaz Serrano, 2017	10	114.68	15.9800 +			114	68	[105.19; 1	25.02]	1.2%	5.2%
Sonnesen, 2007	19	293.18	67.3300	-+	+	293	18	[264.42; 3	325.07]	0.8%	5.2%
Lemos, 2006	36	410.07	52.3200	i i	*	410	07	[393.33; 4	27.52]	5.1%	5.3%
Sonnesen, 2001	26	382.20	63.9000		-#	382	20	[358.41; 4	07.57]	2.1%	5.3%
Sonnesen, 2005	88	362.20	72.7000		*	362	20	[347.32; 3	377.71]	5.0%	5.3%
Araujo, 2016	204	267.60	97.1000	÷		267	60	[254.60; 2	281.26]	3.5%	5.3%
Medhat, 2018	177	107.40	17.5000			107	40	[104.85; 1	10.01]	15.3%	5.3%
Haritha, 2012	40	191.17	11.4700	+		191	17	[187.65; 1	94.76]	25.4%	5.3%
Oueis, 2009	42	604.00	241.8000			<u> </u>	00	[535.13; 6	81.74]	0.6%	5.2%
Kamegai2, 2005	533	393.30	138.5000			393	30	[381.72; 4	05.24]	9.8%	5.3%
Antonarakis, 2015	20	422.80	107.1000	1		422	80	[378.37; 4	72.44]	0.7%	5.2%
Antonarakis, 2012	25	478.80	121.8000			478	80	[433.36; 5	529.01]	0.9%	5.2%
Al Khateeb, 2015	33	325.50	95.2000		•	325	50	[294.59; 3	859.66]	0.9%	5.2%
Common effect mode	I 1923					248	90	[246.57; 2	251.24]	100.0%	
Random effects mode Heterogeneity: $I^2 = 100\%$	$\tau^2 = 0.7$	1906, <i>p</i> =	0			311.	72	[255.99; 3	79.59]		100.0%
5 7.	,	, _F		200 30	0 400 500	600					

Weight Weight Study Mean MLN 95%-CI (fixed) (random) С Kamegai1, 2005 [469.72; 490.92] 41.4% 10.3% 480.20 Kaya, 2017 355.80 [307.98; 411.04] 1.0% 9.8% Thongundomtorn, 2009 [873.19; 961.33] 10.3% 916.20 8.7% Bonjardim, 2005 + [315.57; 336.57] 325.90 19.4% 10.3% 10 Pereira, 2007 326.00 [308.93: 344.01] 7.0% 10.2% Pedroni Pereira, 2018 491.09 [425.19; 567.20] 1.0% 9.8% Kamegai2, 2005 10.3% 462.30 [447.55; 477.53] 19.2% Alkan, 2006 773.00 [666.97; 895.88] 0.9% 9.8% Kampe1, 1987 532.00 0.5% 9.4% [438.57; 645.33] [447.07; 595.56] Kampe2, 1987 516.00 1.0% 9.8% Fixed effect model 456.69 [450.25; 463.21] 100.0% Random effects model 489.35 [399.86; 598.87] 100.0% Heterogeneity: $I^2 = 99\%$, $\tau^2 = 0.1029$, p < 0.01500 -500 0

FIGURE 3. Forest plots comparing the bite force values of all the three dentitions based on the random-effects model. (A) Primary dentition—3 to 6 years, (B) Mixed dentition—6 to 12 years, and (C) Permanent dentition—12 to 19 years. Based on the random-effects model, 3 forest plots were generated for primary, mixed and permanent dentitions. MLN: Log transformed mean; CI: Confidence Interval; SD: Standard Deviation.

Β



FIGURE 4. Forest plots comparing the bite force values between dentition, gender and side. (A) Primary male, (B) Primary female, (C) Primary right, (D) Primary left, (E) Mixed male, (F) Mixed female, (G) Mixed right, (H) Mixed left, (I) Permanent male, (J) Permanent female, (K) Permanent right, (L) Permanent left, (M) Mixed male, (N) Mixed female, (O) Mixed male, (P) Mixed female, (Q) Permanent male, (R) Permanent female, (S) Permanent male, and (T) Permanent female. MLN: Log transformed mean; CI: Confidence Interval; SD: Standard Deviation.

	Study	Total	Mean	SD		Mean		Mean	95%–CI (c	Weight ommon) (r	Weight andom)		Study	Total Mean	SD	Mean	Mean	95%–Cl (c	Weight ommon) (I	Weight random)
Α	sub = Portable OBF ga Kamegai et al. 2005 Hama et al. 2017a Hama et al. 2017b Abu Alhaija et al. 2018 Common effect model Random effects model Heterogeneity: $l^{2} = 8\%$, r^{2}	uge/ Ma 72 1 128 1 134 1085 1 1419	eter 196.00 9 185.11 8 . 9 197.42 6	96.1000 - 39.2300 - 99.6600 35.7400	\$ * III +			196.00 [17] 185.11 [17] 197.42 [19] 196.72 [19] 195.33 [188	5.01; 219.50] 0.28; 201.23] 3.55; 201.37] 3.02; 200.50] 3.44; 202.47]	2.4% 4.4% 0.0% 77.5% 84.3% 	15.9% 16.6% 0.0% 17.4% 49.9%	D	sub = Japan Kamegai et al. 2005 Hama et al. 2017a Hama et al. 2017b Common effect model Random effects model Heterogeneity: $t^2 = 89\%$, t^2	72 196.00 9 128 185.11 8 134 234.14 9 334 = 0.0139, p < 0.	6.1000 9.2300 9.6600	-	196.00 [175.0 185.11 [170.2 234.14 [217.8 208.73 [198.7 204.58 [177.2	; 219.50] 3; 201.23] 5; 251.64] ; 219.24] 5; 236.13]	8.4% 15.4% 20.7% 44.5% 	15.9% 16.7% 17.0% 49.6%
	sub = Pressurized Transducer Rentes et al. 2002 Castelo et al. 2010a Castelo et al. 2010b Common effect model Random effects model Heterogeneity: j ² = 89%, t	30 2 14 2 19 2 63	234.66 4 287.00 5 280.46 4 09. p < 0.	42.9800 50.6300 48.3100		 V	+	234.66 [219 287.00 [26 280.46 [259 260.16 [248 265.57 [233	9.77; 250.56] 1.67; 314.79] 9.56; 303.05] 3.97; 271.87] 3.94; 301.47]	7.1% 3.6% 5.1% 15.7% 	16.9% 16.4% 16.7% 50.1%			30 234.66 4 14 287.00 5 19 280.45 4 63 = 0.0109, p < 0.	2.9800 0.6300 8.3100	+	234.66 [219.7 287.00 [261.6 280.45 [259.5 260.16 [248.90 265.56 [233.94	7; 250.56] 7; 314.79] 6; 303.04] 6; 271.86] 6; 301.46]	25.0% 12.6% 17.9% 55.5% 	17.1% 16.5% 16.8% 50.4%
	Common effect model Random effects model Heterogeneity: <i>I</i> ² = 97%, r Test for subgroup differenc Test for subgroup differenc	1482 ² = 0.034 es (fixed es (rand	47, p < 0. effect): χ om effect	$\chi_1^{01} = 130.677$ (s): $\chi_1^{2} = 20.677$	307 200(22 38, df = 1	2002#10/260 (p < 0.01)	0 280 300	205.56 [202 226.52 [194	2.01; 209.18] 4.47; 263.85]	100.0% 	100.0%		Common effect model Random effects model Heterogeneily: $l^2 = 94\%$, r^2 Test for subgroup difference Test for subgroup difference	397 = 0.0297, p < 0. as (fixed effect): y as (random effect	$\sum_{1}^{2} = 42.83160-200x$ s): $\chi_{1}^{2} = 7.14$, df = 1	220.0240 260 280 300 (p < 0.01)	235.88 [228.2] 233.38 [202.4]	; 243.74] ; 269.08]	100.0%	 100.0%
	Study	Tota	al Mea	n Si	D	Me	an	Mean	95%-C	Weigh I (common	t Weight) (random)		Study	Total Mean	SD	Mean	Mean	95%–CI	Weight common)	Weight (random)
В	sub = Pressurized rubi Sonnesen et al. 2001 Gaviao et al. 2005 Gaviao et al. 2006 Sonnesen et al. 2007 Lemos et al. 2007 Common effect model Dendem difect model	2 2 2 2 2 2 2 2 2 2 2 2 2 2 2 2 2 2 2	6 382.2 8 362.2 5 441.6 9 293.1 6 410.0 2 302.0	0 63.900 0 72.700 3 44.300 8 67.330 7 52.320 0 61.000	0 0 0 0 0 0		*	382.20 362.20 441.63 293.18 410.07 302.00 383.75	[358.41; 407.57 [347.32; 377.71 [419.77; 464.63 [264.42; 325.07 [393.33; 427.52 [277.56; 328.60 [375.28; 392.41] 2.89] 6.79] 4.69] 1.19] 6.89] 1.69] 23.69	6 7.1% 6 7.2% 6 7.2% 6 7.1% 6 7.2% 6 7.1% 6 7.1%	Е	sub = Denmark Sonnesen et al. 2001 Sonnesen et al. 2005 Sonnensen et al. 2007 Common effect model Random effects model Heterogeneity: <i>I²</i> = 89%, τ ²	26 382.20 88 362.20 19 293.18 133 = 0.0168, p < 0.	63.9000 72.7000 67.3300	*	382.20 [358. 362.20 [347. 293.18 [264. 359.48 [347.] 345.49 [296.]	41; 407.57] 32; 377.71] 42; 325.07] '3; 371.64] 51; 402.42]	5.6% 13.1% 2.2% 20.8% 	10.0% 10.0% 9.9% 29.9%
	Heterogeneity: <i>I</i> ² = 95%, τ sub = Portable OBF ga Kamegai et al. 2005a Kamegai et al. 2005a Gudapaneni RK et al, 20 Haritha et al. 2012 Antonarakis et al. 2015	² = 0.025 uge/ me 48 53 020 12 4 2 3	51, <i>p</i> < 0. eter 16 296.3 13 393.3 10 320.9 10 191.1 10 422.8 13 325 5	01 0 128.100 0 138.500 3 54.230 7 11.470 0 107.100 0 95.200	0 0 0		*	296.30 393.30 320.93 191.17 	[285.13; 307.91 [381.72; 405.24 [311.37; 330.76 [187.65; 194.76 [378.37; 472.64 [378.37; 472.64]] 8.0%] 13.1%] 12.9%] 34.0%] 1.0%	6 7.2% 6 7.2% 6 7.2% 6 7.2% 6 7.2% 6 7.0% 6 7.0%		Gaviao et al. 2006 Lemos et al. 2006 Serra et al. 2007 Araujo et al. 2016 Diaz Serrano et al. 2017 Common effect model Random effects model Heterogeneity: / ² = 100%, ·	15 441.63 36 410.07 22 302.00 204 267.60 10 114.68 287 $c^2 = 0.2875, p < 0$	44.3000 52.3200 61.0000 97.1000 15.9800 + .01	*	 441.63 [419. 410.07 [393. 302.00 [277. 267.60 [254. 114.68 [105. 329.89 [321.1 278.81 [174. 	77; 464.63] 33; 427.52] 56; 328.60] 50; 281.26] 19; 125.02] 15; 338.13] 10; 446.50]	8.9% 13.3% 9.3% 3.1% 37.8%	10.0% 10.0% 9.9% 10.0% 9.9%
	Al Knateeb et al. 2015 Common effect model Random effects model Heterogeneity: $I^2 = 100\%$, sub = Dynamometer	123 τ ² = 0.07	13 325.5 12 776, p = (0 95.200	U	*	-	325.50 258.02 314.71	[254.59; 359.66 [254.70; 261.38 [251.42; 393.94] 70.19] 70.19	6 7.1% 6 - 42.9%		sub = Japan Kamegai et al. 2005a Kamegai et al. 2005b Common effect model Random effects model	486 296.30 1 533 393.30 1 1019	28.1000 38.5000	*	296.30 [285. 393.30 [381. 353.48 [345. 341.46 [258.]	13; 307.91] 72; 405.24] 24; 361.92] 71: 450.69]	15.6% 25.8% 41.4%	10.0% 10.1% 20.1%
	Araujo et al. 2016 Diaz Serrano et al. 2017 Common effect model Random effects model Heterogeneity: $I^2 = 100\%$,	$\tau^2 = 0.35$	4 267.6 0 114.6 4 577, p < 0	0 97.100 8 15.980 	0 +	*		267.60 114.68 216.61 175.32	[254.60; 281.26 [105.19; 125.02 [207.47; 226.16 [76.42; 402.21] 4.7%] 1.6%] 6.3%]	6 7.2% 6 7.1% 6 - 14.3%		Heterogeneity: $l^2 = 99\%$, τ^2 Common effect model Random effects model Heterogeneity: $l^2 = 99\%$, τ^2 Test for subgroup difference	 = 0.0398, ρ < 0. 1439 = 0.1473, ρ < 0. as (fixed effect) as a start of the start o	01 01 2-22.55 df 4-20(9)	200250 300 350 400 4	345.58 [340.3 309.32 [243.0	7; 350.86] 4; 392.71]	100.0% 	100.0%
	Common effect model Random effects model Heterogeneity: $l^2 = 100\%$, Test for subgroup difference Test for subgroup difference	165 $\tau^2 = 0.12$ es (fixed es (rand	i2 273, p = 0 effect): γ om effect	$\chi_2^2 = 1055.7^{-1}$ (s): $\chi_2^2 = 3.7^{-1}$	I, df =1 50 (,4 6, df = 2 ()	20002830 3 p = 0.15)	00 350 400	280.25 307.37 450	[277.23; 283.30 [254.76; 370.85] 100.0%]	6 - 100.0%		Test for subgroup difference	(random effect	$\chi_2^2 = 0.72$, df = 2	(p = 0.70)				
	Study	Total	Mean	SD		Mean		Mean	95%-CI (Weight common) (I	Weight random)		Study	Total Mean	SD	Mean	Mean	95%–CI (Weight common) (Weight random)
С	sub = Portable OBF ga Kamegai et al. 2005a Kamegai et al. 2005b Common effects model Random effects model Heterogeneity: $I^2 = 72\%$, t	uge/me 902 4 600 4 1502 ² = 0.000	ter 480.20 1 462.30 1 05, p = 0.	162.4000 187.3000		■ +		480.20 [44 462.30 [44 474.46 [46 472.08 [45	69.72; 490.92] 47.55; 477.53] 55.88; 483.19] 54.91; 489.89]	46.8% 21.7% 68.5% 	17.5% 17.4% 34.9%	F	sub = Japan Kamegai et al. 2005a Kamegai et al. 2005b Common effect model Random effects model Heterogeneity: $J^2 = 72\%, \tau^2$	902 480.20 1 600 462.30 1 1502 = 0.0005, p = 0.	62.4000 87.3000	≡ +	480.20 [469. 462.30 [447. 474.46 [465.8 472.08 [454.9	72; 490.92] 55; 477.53] 8; 483.19] 11; 489.89]	46.8% 21.7% 68.5% 	17.5% 17.4% 34.9%
	sub = Bite Force Record Kampe et al. 1987a Kampe et al. 1987b Common effect model Random effects model Heterogeneity: $I^2 = 0\%$, τ^2	rder 13 5 16 5 29 = 0, p =	532.00 1 516.00 1 0.80	189.0000 151.0000		V V		- 532.00 [4: 516.00 [4: 521.63 [4: 521.63 [4:	38.57; 645.33] 47.07; 595.56] 54.91; 585.28] 54.91; 585.28]	0.6% 1.1% 1.7%	14.6% 15.8% 30.3%		sub = Sweden Kampe et al. 1987a Kampe et al. 1987b Common effect model Random effects model Heterogeneity: $I^2 = 0\%$, τ^2 :	13 532.00 1 16 516.00 1 29 = 0, p = 0.80	89.0000 51.0000		- 532.00 [438. 516.00 [447. 521.63 [464.9 521.63 [464.9	57; 645.33])7; 595.56] 11; 585.28] 11; 585.28]	0.6% 1.1% 1.7% 	14.6% 15.8% 30.3%
	sub = Pressurized Tran Bonjardim et al. 2005 Pereira et al. 2007 Common effect model Random effects model Heterogeneity: $J^2 = 0\%$, τ^2	20 3 20 3 40	325.90 326.00	23.9600 40.0000 -	# ★- ◇ ◇	************		325.90 [3 326.00 [30 325.93 [31 325.93 [31	15.57; 336.57] 08.93; 344.01] 17.04; 335.06] 17.04; 335.06]	21.9% 7.9% 29.8% 	17.4% 17.3% 34.7%		sub = Brazil Bonjardim et al. 2005 Pereira et al. 2007 Common effect model Random effects model Heterogeneity: $I^2 = 0\%, \tau^2$	20 325.90 20 326.00 40	23.9600 = 40.0000 - ¢ ¢		325.90 [315. 326.00 [308. 325.93 [317.0 325.93 [317.0	57; 336.57] 33; 344.01] 44; 335.06] 44; 335.06]	21.9% 7.9% 29.8% 	17.4% 17.3% 34.7%
	Common effect model Random effects model Heterogeneity: $J^2 = 99\%$, τ Test for subgroup difference Test for subgroup difference	1571 ² = 0.047 es (fixed es (rand	75, p < 0. effect): γ	01 $\chi_2^2 = 506.28$, is): $\chi_2^2 = 278$	d8502400	9: 450)500 2 (p < 0.01)) 550 600	424.88 [41 428.08 [35	18.51; 431.34] 57.98; 511.91]	100.0% 	 100.0%		Common effect model Random effects model Heterogeneity: $l^2 = 99\%$, τ^2 Test for subgroup difference Test for subgroup difference	1571 = 0.0475, <i>p</i> < 0. es (fixed effect): χ es (random effect	$\chi^2_2 = 506.28, d s 502.4$ s): $\chi^2_2 = 278.77, d f =$	06: 450)500 550 600 2 (p < 0.01)	424.88 [418.9 428.08 [357.9	i1; 431.34] 18; 511.91]	100.0% 	 100.0%

FIGURE 5. Forest plots comparing the bite force values between site of measurement, device and ethnicity for each included dentition. (A–C), Site of measurement and device, where A corresponds to Primary dentition, B—mixed dentition and C—Permanent Dentition. (D–F), Site of measurement and ethnicity, where D corresponds to Primary dentition, E—mixed dentition and F—Permanent Dentition. MLN: Log transformed mean; CI: Confidence Interval; SD: Standard Deviation.

among females (n = 106) computed to 23.61 (20.87; 26.70) kg. MVBF of males (n = 60) measured in kgf averaged 30.29 (22.74; 40.34), and that among females (n = 80) measured in kgf averaged 23.43 (18.78; 29.24).

Six studies tested BF using a pressurized rubber tube in the mixed dentition at the permanent 1st molar site which equated to 362.64 (318.55; 412.84) N while portable OBF gauges averaged 314.71 (251.42; 393.94) N from six included studies. The dynamometer studies by Araujo *et al.* [52], 2016 and Diaz Serrano *et al.* [65] 2017 resulted in an average BF of 175.32 (76.42; 402.21) N (Fig. 5B).

According to ethnicity, the highest number of studies were reported from Brazil (n = 5) resulting in an average BF of 278.81 (174.10; 446.50) N, followed by Denmark (n = 3)

and Japan (n = 2)—which reported average BF values of 345.49 (296.61; 402.42) N and 341.46 (258.71; 450.69) N respectively. (Fig. 5E).

3.3.3 Permanent dentition (ages 12 to 19 years)

From ten included studies and 2673 participants, the average BF of children with only permanent teeth was estimated as 489.35 (399.86; 598.87) N. Studies revealed that males (n = 792) had a greater masticatory force of 547.57 (537.11; 558.24) N than females (n = 898) with 424.71 (416.48; 433.10) N. Molar BF of the right side computed to 441.59 (315.77; 617.53) N and that on the left to 411.15 (281.54; 600.44) N. Similar observations were made while comparing bite forces

of males (n = 305) 40.12 (29.62; 54.36) and females (n = 327) 31.96 (20.06; 50.90) in kg, and of males (n = 60) 40.51 (35.29; 46.51) and females (n = 80) 36.13 (28.55; 45.74) in kgf.

A total of six out of the 18 included studies contributed to the analyses of parameters, specific to the site (permanent 1st molar), device, and ethnicity. The values for device and ethnicity reported similar BF estimates as given in (Fig. 5C and Fig. 5F).

3.3.4 Overlapping dentition (combination of ages from 3 through 19)

A few authors reported studies in which participants (n = 1102) had overlapping dentitions (*i.e.*, primary and mixed together or mixed and permanent together) with various units for analysis. Meta-analysis could not be performed since the collected data from the above studies were diverse due to the lack of a clear distinction of the recorded dentition. However, one study [32] showed that BF measured in kgf differed based on gender when subjected to meta-analysis. The males (n = 356) and females (n = 322) of the study had BF values of 40.51 (35.29; 46.51) kgf and 36.13 (28.55; 45.74) kgf, respectively.

3.4 Review of clinical trials

Subsequent to the consideration of heterogeneity in the included clinical trials, we qualitatively summarized the pre and post-intervention results.

Observations from the qualitative description of the ten included clinical trials revealed various treatment modalities, such as orthodontic management of malocclusion (n = 5), while two studies demonstrated an increase in the BF values post-treatment with orthodontic appliances [33, 34]. Two studies [35, 36] reported a decrease in the BF estimation, post-treatment. The results reported an immediate decrease post-treatment with an expansion plate and quad helix while data showed an increased BF value at a 4–6-month follow-up using the same appliance [37].

Restorative management for caries-affected teeth (n = 2) included stainless steel crowns [38] and conventional glassionomer restorations [18]. Both studies showed a clinically significant increase in MVBF results post-treatment.

Prosthetic replacement of missing teeth with a removable partial denture (n = 1) showed an increase in the masticatory force at 6–12 month follow-up [39]. A steady increase in MVBF was observed in a report by Alkan *et al.* [40] (2006) post-surgical osteotomy in children with mandibular deficiency, at five intervals. Also, a trial by Martini *et al.* [13] which included healthy controls and a study group with craniosynostosis, showed that the MVBF values of the healthy controls were greater than that of the study group.

4. Discussion

This meta-analysis aimed to quantify the MVBF of children based on age, gender, site, ethnicity, measuring device and side for this systematic review. While several studies have successfully calibrated BF among various types of dentitions, to the best of our knowledge, this is the first systematic review on pooled MVBF values in children and adolescents based on age and dentition. The principal findings of this study showed the magnitude of BF to be comparatively greater in the permanent dentition than in the mixed and primary dentition. Also, results observed in the 'gender' group concerning age were correlative. Males and females with permanent dentition had greater BF than those with mixed and primary dentitions. A study by Usui *et al.* [32] (2007) reported that BF tended to increase with age, up to 20 years in males and 17 years in females. Kamegai *et al.* [41] (2005) found that mean BF increased in females until age 14. A gradual increase was noted thereafter until age 17; however, the mean value did not exceed the mean recorded at age 14.

The secondary findings showed a significant association between gender and MVBF. Males had greater BF than females in all types of dentitions, an observation consistent with the interpretations of Abu-Alhaija et al. [30] (2018), Palinkas et al. [42] (2010), Ingervall and Minder [43] (1997), and Shiau and Wang [44] (1993). This could be because males had more confidence [15] when recording BF and their anatomic [36] and hormonal variations could have contributed to this increase [45]. Also, a greater dental size concomitant with a larger periodontal ligament area [10] results in males developing a more pronounced masticatory ability. In contrast, several authors have reported no gender influence on occlusal forces other than elements such as age, body weight, and height [46-50]. The right side showed greater masticatory forces than the left in all three dentitions. Though multiple studies [15, 28– 31, 38, 65, 84] have reported the MVBF values estimated bilaterally in their respective subjects, the rationale for the same is unclear. 'Side dominance' (right and left-handedness) could be one possible reason for the increase in right-side BF value and further studies are required to understand the factors influencing the changes in the measured values. The results of this study based on side, showed a decline in the MVBF values of the mixed dentition while comparing the primary dentition. A possible explanation for this underestimation could be due to included studies recruiting children with mixed dentition but conducted the analysis on either permanent first molars or primary molars. This combination of sites for analysis could have altered the magnitude of the biting force. Further, the inclusion of children in various stages of dental development, such as eruption and incipient exfoliation, might have led to the improper recording of occlusal pressure. Nevertheless, a comparison between primary and permanent dentitions showed approximately a two-fold increase in the permanent dentition values for side.

The present systematic review reports 11 different types of devices for recording BF and the data analysis is presented in **Supplementary Table 1**. Among these, the portable occlusal bite force gauge/meter (OBF/OBG) is the most commonly used due to its portability, comfortable biting element, and instant digital display of BF. Exclusive to the primary dentition, the pressurized transducer and dental pre-scale systems are pre-ferred devices because the vinyl rubber tube is confined to the occlusal pattern, facilitating accurate reflection of BF.

Most of the studies included were cross-sectional, which affected the generalizability of the results. Harmonization of study designs can compensate for any dissimilarity caused by including observational studies and trials. Second, it would be helpful for future research if inter- and intra-comparison of the most influential variables, such as age, dentition, gender, side, and site, could be assessed since a selective omission in a few of the included studies can influence the outcome. For variables such as height [31, 51], weight [16, 31], and BMI [9, 15–17, 52, 53] a clear correlation among included studies could not be established since a few authors reported a direct association while a few contradicted the same. This conflict could be due to recruitment bias. Therefore, study participants should be equally identified across the spectrum of obese, overweight, healthy, underweight, and malnourished individuals. Also, recording the status of the dentition is essential, since the presence of caries [54] and restorations [14, 55] alters the structural integrity and pain threshold of the tooth, thereby reducing the masticatory function. Moreover, the absence of teeth [31] leads to reduced BF due to loss of antagonistic contacts. Malocclusions such as aberrant overjet [30] overbite [31], and cross-bite [56–58] can affect the child's ability to bite since fewer occlusal contact points result in weaker occlusal support and musculature [8]. Therefore, further studies are required to understand the relation between occlusion, tooth size, and BF. The current findings call attention to the fact that, among the world regions included, there is a lack of significant data on MVBF from certain areas of the globe, such as the continents of Australia and Africa. The occlusal table for the Australian population is more pronounced mesiodistally [59], which may contribute to a higher value of MVBF. Thus, consideration of body and dental health variables is required for comprehensible data and improved study quality.

Third, few authors have reported findings as mean average values only and lack of standard deviation resulted in the exclusion of studies from the meta-analysis. An understanding of these drawbacks in published studies and the design of future clinical trials aimed at recording MVBF in healthy versus affected individuals with a uniform reporting protocol is required to better understand the developmental and environmental influences over time. Few authors assessed BF using a custom-made device. However, the reliability and reproducibility of such self-made equipment are questionable. Therefore, standardized devices with established sensitivity and specificity should be considered. Apart from longitudinal studies targeted at specific dentitions to assess the role of factors influencing BF, studies to enable standardization of BF recording devices are required to rule out methodological biases.

This systematic review had both strengths and limitations. Our current findings will likely provide dental professionals with the highest possible available evidence on MVBF including various devices and sites to assess the same. Therefore, the concept of BF, if applied to regular clinical practice with the results of this study, can prove to be a useful adjunct investigative tool for the diagnosis of any deviation from normal function and development. This review provides the first comprehensive analysis of mean BF values in all types of dentitions-primary, early to late mixed, permanent dentition and, combination dentition among children and adolescents. Our findings revealed average BF values range from 246.22–489.35 N and 5.69–16.1 kg in children and adolescents. In addition, the average BF values for devices ranged from 226.52 to

428.08 N. The results of this review provide clinicians with an insight on the amount of biting force that can exist in each dentition and may serve as a baseline value for future studies and clinical assessment. The absence of time limits for the inclusion of studies available in the literature renders the evidence presented robust. Multiple factors influencing BF have been addressed in this review, age being the most vital variable [60], along with a direct method of BF assessment. Subgroup analysis included other variables such as gender and side. Other factors influencing BF and the accuracy of the apparatus used for recording BF were not considered in this review. Also, we deviated from the registered protocol because we were unable to assess the pre and post-MVBF values of the various clinical trials. The lack of information (raw data) from 81 studies, despite multiple reminders, could also have affected the results.

5. Conclusion

1. Systematic analysis of the data showed that BF ranged from 246.22–489.35 N and 5.69–16.1 kg in children and adolescents and the average BF values for devices ranged from 226.52 to 428.08 N.

2. The portable occlusal bite force gauge/meter (OBF/OBG) was the most commonly used device to record BF in all dentitions.

3. The transition from primary to early-mixed to late-mixed and permanent dentitions is a dynamic process. Hence, the BF taken at a given period may differ from one type of dentition to another.

4. Not only physiologic changes but also pathologic conditions like caries, malocclusion, and the early shedding of primary teeth can affect BF values.

5. Thus, BF is an entity that needs close and continuous monitoring for clinical relevance.

AVAILABILITY OF DATA AND MATERIALS

The data are contained within this article and supplementary material.

AUTHOR CONTRIBUTIONS

PJ, GFP and MSM—Contributed to conception, design, data acquisition, analysis and interpretation, drafted and critically revised the manuscript. RK—Contributed to analysis and interpretation and critically revised the manuscript. NP and SSA—Contributed to data acquisition and critically revised the manuscript. All authors contributed to editorial changes in the manuscript. All authors read and approved the final manuscript.

ETHICS APPROVAL AND CONSENT TO PARTICIPATE

Not applicable.

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CONFLICT OF INTEREST

The authors declare no conflict of interest.

SUPPLEMENTARY MATERIAL

Supplementary material associated with this article can be found, in the online version, at https://oss.jocpd.com/ files/article/1651170756277485568/attachment/ Supplementary%20material.docx.

REFERENCES

- [1] Calderon Pdos S, Kogawa EM, Lauris JR, Conti PC. The influence of gender and bruxism on the human maximum bite force. Journal of Applied Oral Science. 2006; 14: 448–453.
- [2] Alhowaish L. Bite force evaluation in children following dental treatment [Doctoral Thesis]. England: University of Leeds. 2012.
- [3] Gallagher S, O'Connell BC, O'Connell AC. Assessment of occlusion after placement of stainless steel crowns in children—a pilot study. Journal of Oral Rehabilitation. 2014; 41: 730–736.
- [4] Alshareef AA, Alkhuriaf A, Pani SC. An evaluation of bite pattern in children with severe-early childhood caries before and after complete dental rehabilitation. Pediatric Dentistry. 2017; 39: 455–459.
- [5] Testa M, Geri T, Pitance L, Lentz P, Gizzi L, Erlenwein J, et al. Alterations in jaw clenching force control in people with myogenic temporomandibular disorders. Journal of Electromyography and Kinesiology. 2018; 43: 111–117.
- [6] Zwir L, Fraga M, Sanches M, Hoyuela C, Len C, Terreri MT. A casecontrol study about bite force, symptoms and signs of temporomandibular disorders in patients with idiopathic musculoskeletal pain syndromes. Advances in Rheumatology. 2018; 58: 7.
- [7] Makino E, Nomura M, Motegi E, Iijima Y, Ishii T, Koizumi Y, *et al.* Effect of orthodontic treatment on occlusal condition and masticatory function. The Bulletin of Tokyo Dental College. 2014; 55: 185–197.
- [8] Roldán SI, Restrepo LG, Isaza JF, Vélez LG, Buschang PH. Are maximum bite forces of subjects 7 to 17 years of age related to malocclusion? The Angle Orthodontist. 2016; 86: 456–461.
- ^[9] Sun KT, Chen SC, Li YF, Chiang HH, Tsai HH, Li CY, *et al.* Biteforce difference among obese adolescents in central Taiwan. Journal of the Formosan Medical Association. 2016; 115: 404–410.
- [10] Ferrario VF, Sforza C, Zanotti G, Tartaglia GM. Maximal bite force in healthy young adults as predicted by surface electromyography. Journal of Dentistry. 2004; 32: 451–457.
- [11] Koc D, Dogan A, Bek B. force and influential factors on bite force measurements: a literature review. European Journal of Dentistry. 2010; 4: 223–232.
- ^[12] Ohira A, Ono Y, Yano N, Takagi Y. Effect of chewing exercise in preschool children on maximum bite force and masticatory performance. International Journal of Paediatric Dentistry. 2012; 22: 146–153.
- ^[13] Martini M, Wiedemeyer V, Heim N, Messing-Jünger M, Linsen S. Bite force and EMG evaluation after cranioplasty in patients with craniosynostosis. Oral Surgery, Oral Medicine, Oral Pathology and Oral Radiology. 2017; 124: e267–e275.
- [14] Kampe T, Haraldson T, Hannerz H, Carlsson GE. Occlusal perception and bite force in young subjects with and without dental fillings. Acta Odontologica Scandinavica. 1987; 45: 101–107.

- ^[15] Varga S, Spalj S, Lapter Varga M, Anic Milosevic S, Mestrovic S, Slaj M. Maximum voluntary molar bite force in subjects with normal occlusion. European Journal of Orthodontics. 2011; 33: 427–433.
- [16] Gavião MBD, Raymundo VG, Rentes AM. Masticatory performance and bite force in children with primary dentition. Brazilian Oral Research. 2007; 21: 146–152.
- [17] Su CM, Yang YH, Hsieh TY. Relationship between oral status and maximum occlusal bite force in preschool children. Journal of Dental Sciences. 2009; 4: 32–39.
- [18] Subramaniam P, Girish Babu KL, Ifzah. Effect of restoring carious teeth on occlusal bite force in children. Journal of Clinical Pediatric Dentistry. 2016; 40: 297–300.
- [19] Verma TP, Kumathalli KI, Jain V, Kumar R. Bite force recording devices—a review. Journal of Clinical and Diagnostic Research. 2017; 11: ZE01–ZE05.
- [20] Page MJ, McKenzie JE, Bossuyt PM, Boutron I, Hoffmann TC, Mulrow CD, *et al.* The PRISMA 2020 statement: an updated guideline for reporting systematic reviews. BMJ. 2021; 372: n71.
- [21] Adolescent health/Health topics by WHO (World Health Organization). 2019. Available at: https://apps.who.int/adolescent/seconddecade/section2/page1/recognizing-adolescence.html (Accessed: 30 August 2019).
- ^[22] Wells G, Shea B, O'Connell D, Robertson J, Peterson J, Welch V, *et al.* the newcastle-ottawa scale (NOS) for assessing the quality of non-randomized studies in meta-analyses. 3rd Symposium on Systematic Reviews: Beyondthe Basics. Improving Quality and Impact, Oxford, UK. 2013.
- [23] Modesti PA, Reboldi G, Cappuccio FP, Agyemang C, Remuzzi G, Rapi S, *et al.* Panethnic differences in blood pressure in Europe: a systematic review and meta-analysis. PLoS One. 2016; 11: e0147601.
- [24] Sterne JA, Hernán MA, Reeves BC, Savović J, Berkman ND, Viswanathan M, *et al.* ROBINS-I: a tool for assessing risk of bias in nonrandomised studies of interventions. BMJ. 2016; 355: i4919.
- [25] Viechtbauer W. Conducting meta-analyses in R with the metaphor package. Journal of Statistical Software. 2010; 36: 1–48.
- [26] Higgins JP, Thompson SG, Deeks JJ, Altman DG. Measuring inconsistency in meta-analyses. BMJ. 2003; 327: 557–560.
- ^[27] Kiriishi K, Doi H, Magata N, Torisu T, Tanaka M, Ohkubo M, *et al.* Occlusal force predicts global motion coherence threshold in adolescent boys. BMC Pediatrics. 2018; 18: 331.
- [28] Karibe H, Ogata K, Hasegawa Y, Ogihara K. Relation between clenching strength and occlusal force distribution in primary dentition. Journal of Oral Rehabilitation. 2003; 30: 307–311.
- [29] Mountain G, Wood D, Toumba J. Bite force measurement in children with primary dentition. International Journal of Paediatric Dentistry. 2011; 21: 112–118.
- [30] Abu-Alhaija E, Owais AI, Obaid H. Maximum occlusal bite force in preschool children with different occlusal patterns. Journal of Clinical and Experimental Dentistry. 2018; 10: e1063–e1068.
- [31] Heydari A, Nakhjavani YB, Anaraki EA, Arvan S, Shafizadeh M. Bite force of 3–6-year-old children after unilateral extraction of primary teeth. Journal of Dentistry. 2018; 15: 47–53.
- [32] Usui T, Uematsu S, Kanegae H, Morimoto T, Kurihara S. Change in maximum occlusal force in association with maxillofacial growth. Orthodontics & Craniofacial Research. 2007; 10: 226–234.
- [33] Varga S, Spalj S, Anic Milosevic S, Lapter Varga M, Mestrovic S, Trinajstic Zrinski M, *et al.* Changes of bite force and occlusal contacts in the retention phase of orthodontic treatment: a controlled clinical trial. American Journal of Orthodontics and Dentofacial Orthopedics. 2017; 152: 767–777.
- [34] Antonarakis GS, Kiliaridis S. Predictive value of masseter muscle thickness and bite force on Class II functional appliance treatment: a prospective controlled study. European Journal of Orthodontics. 2015; 37: 570–577.
- [35] Antonarakis GS, Kjellberg H, Kiliaridis S. Predictive value of molar bite force on Class II functional appliance treatment outcomes. European Journal of Orthodontics. 2012; 34: 244–249.
- [36] Al-Khateeb SN, Abu Alhaija ES, Majzoub S. Occlusal bite force change after orthodontic treatment with Andresen functional appliance. European Journal of Orthodontics. 2015; 37: 142–146.

- [37] Sonnesen L, Bakke M. Bite force in children with unilateral crossbite before and after orthodontic treatment. A prospective longitudinal study. European Journal of Orthodontics. 2007; 29: 310–313.
- [38] Owais AI, Al-Battah AH, Abu Alhaija ES. Changes in occlusal bite force following placement of preformed metal crowns on primary molars in 4–6 years old children: a 6 months' follow-up pilot study. European Archives of Paediatric Dentistry. 2019; 20: 9–14.
- [39] Serra MD, Gambareli FR, Gavião MB. A 1-year intraindividual evaluation of maximum bite force in children wearing a removable partial dental prosthesis. Journal of Dentistry for Children. 2007; 74: 171–176.
- [40] Alkan A, Arici S, Sato S. Bite force and occlusal contact area changes following mandibular widening using distraction osteogenesis. Oral Surgery, Oral Medicine, Oral Pathology, Oral Radiology. 2006; 101: 432–436.
- [41] Kamegai T, Tatsuki T, Nagano H, Mitsuhashi H, Kumeta J, Tatsuki Y, et al. A determination of bite force in northern Japanese children. European Journal of Orthodontics. 2005; 27: 53–57.
- [42] Palinkas M, Nassar MSP, Cecílio FA, Siéssere S, Semprini M, Machadode-Sousa JP, *et al.* Age and gender influence on maximal bite force and masticatory muscles thickness. Archives of Oral Biology. 2010; 55: 797– 802.
- [43] Ingervall B, Minder C. Correlation between maximum bite force and facial morphology in children. The Angle Orthodontist. 1997; 67: 415–424.
- [44] Shiau YY, Wang JS. The effects of dental condition on hand strength and maximum bite force. Cranio. 1993; 11: 48–54.
- [45] Pizolato RA, Gavião MBD, Berretin-Felix G, Sampaio ACM, Trindade Junior AS. Maximal bite force in young adults with temporomandibular disorders and bruxism. Brazilian Oral Research. 2007; 21: 278–283.
- [46] Linderholm H, Lindqvist B, Ringqvist M, Wennström A. Isometric bite force in children and its relation to body build and general muscle force. Acta Odontologica Scandinavica. 1971; 29: 563–568.
- [47] Lindqvist B, Ringqvist M. Bite force in children with bruxism. Acta Odontologica Scandinavica. 1973; 31: 255–259.
- [48] Graber TM. Maximum bite force values of children in different age groups. American Journal of Orthodontics. 1984; 86: 170.
- [49] Ranta R, Tulensalo T, Helle A. Individual changes of maximum bite force in children. IRCS Medical Science. 1984; 12: 1020.
- [50] Kiliaridis S. Masticatory muscle influence on craniofacial growth. Acta Odontologica Scandinavica. 1995; 53: 196–202.
- [51] Duarte Gavião MB, Durval Lemos A, Diaz Serra M, Riqueto Gambareli F, Nobre Dos Santos M. Masticatory performance and bite force in relation to signs and symptoms of temporomandibular disorders in children. Minerva Stomatologica. 2006; 55: 529–539.
- [52] Araujo DS, Marquezin MCS, Barbosa TDS, Gavião MBD, Castelo PM. Evaluation of masticatory parameters in overweight and obese children. European Journal of Orthodontics. 2016; 38: 393–397.
- [53] Medhat AH, Al-Haidar AH. Maximum bite force among Iraqi primary school children in mixed dentition. Indian Journal of Public Health Research & Development. 2019; 10: 800–805.
- [54] Kaya MS, Akyuz S, Guclu B, Diracoglu D, Yarat A. Masticatory parameters of children with and without clinically diagnosed caries in permanent dentition. European journal of Paediatric Dentistry. 2017; 18: 116–120.
- [55] Awawdeh L, Hemaidat K, Al-Omari W. Higher maximal occlusal bite force in endodontically treated teeth versus vital contralateral counterparts. Journal of Endodontics. 2017; 43: 871–875.
- [56] Sonnesen L, Bakke M, Solow B. Bite force in pre-orthodontic children with unilateral crossbite. European Journal of Orthodontics. 2001; 23: 741–749.
- [57] Castelo PM, Gavião MBD, Pereira LJ, Bonjardim LR. Maximal bite force, facial morphology and sucking habits in young children with functional posterior crossbite. Journal of Applied Oral Science. 2010; 18: 143–148.
- [58] Marquezin MCS, Gauch CG, De Siqueira CA, Kobayashi FY, Fonseca FLA, Castelo PM. Evaluation of masticatory and salivary parameters in preschool children with different morphological occlusion. Brazilian Dental Science. 2017; 20: 38–46.
- [59] Sujitha P, Bhavyaa R, Muthu MS, Kirthiga M. Morphological variations and prevalence of aberrant traits of primary molars. Annals of Human

Biology. 2021; 48: 294-306.

- [60] Goodarzi F, Mahvi AH, Hosseini M, Nodehi RN, Kharazifard MJ, Parvizishad M. Prevalence of dental caries and fluoride concentration of drinking water: a systematic review. Dental Research Journal. 2017; 14: 163–168.
- [61] Al Qassar SSS, Mavragani M, Psarras V, Halazonetis DJ. The anterior component of occlusal force revisited: direct measurement and theoretical considerations. European Journal of Orthodontics. 2016; 38: 190– 196.
- [62] Alam MK, Alfawzan AA. Maximum voluntary molar bite force in subjects with malocclusion: multifactor analysis. Journal of International Medical Research. 2020; 48: 030006052096294.
- [63] Bonjardim LR, Gavião MB, Pereira LJ, Castelo PM. Bite force determination in adolescents with and without temporomandibular dysfunction. Journal of Oral Rehabilitation. 2005; 32: 577–583.
- [64] Castelo PM, Pereira LJ, Bonjardim LR, Gavião MBD. Changes in bite force, masticatory muscle thickness, and facial morphology between primary and mixed dentition in preschool children with normal occlusion. Annals of Anatomy. 2010; 192: 23–26.
- [65] Díaz-Serrano KV, Dias TM, Vasconcelos P, Sousa LG, Siéssere S, Regalo S, *et al.* Impact of temporomandibular disorders on the stomatognathic system in children. Medicina Oral Patología Oral Y Cirugia Bucal. 2017; 12: e723–e729.
- [66] Ferreira B, Da Silva GP, Gonçalves CR, Arnoni VW, Siéssere S, Semprini M, et al. Stomatognathic function in Duchenne muscular dystrophy: a case-control study. Developmental Medicine & Child Neurology. 2016; 58: 516–521.
- [67] Fields HW, Proffit WR, Case JC, Vig KWL. Variables affecting measurements of vertical occlusal force. Journal of Dental Research. 1986; 65: 135–138.
- [68] Hama Y, Hosoda A, Komagamine Y, Gotoh S, Kubota C, Kanazawa M, et al. Masticatory performance-related factors in preschool children: establishing a method to assess masticatory performance in preschool children using colour-changeable chewing gum. Journal of Oral Rehabilitation. 2017; 44: 948–956.
- [69] Sathyanarayana HP, Premkumar S. Assessment of maximum voluntary bite force in children and adults with normal occlusion. International Journal of Pharmacutical Science and Health Care. 2012; 1: 64–70.
- [70] Jeong CW, Kim KH, Jang HW, Kim HS, Huh JK. The relationship between oral tori and bite force. Cranio. 2019; 37: 246–253.
- [71] Lemos AD, Gambareli FR, Serra MD, Pocztaruk RDL, Gavião MBD. Chewing performance and bite force in children. Brazilian Journal of Oral Sciences. 2006; 5: 1101–1108.
- [72] Maki K, Nishioka T, Morimoto A, Naito M, Kimura M. A study on the measurement of occlusal force and masticatory efficiency in school age Japanese children. International Journal of Paediatric Dentistry. 2001; 11: 281–285.
- [73] Marquezin MCS, Pedroni-Pereira A, Araujo DS, Rosar JV, Barbosa TS, Castelo PM. Descriptive analysis of the masticatory and salivary functions and gustatory sensitivity in healthy children. Acta Odontologica Scandinavica. 2016; 74: 443–448.
- [74] Oueis H. Factors affecting masticatory performance of Japanese children. International Journal of Paediatric Dentistry. 2009; 19: 201–205.
- [75] Pedroni-Pereira A, Marquezin MCS, Araujo DS, Pereira LJ, Bommarito S, Castelo PM. Lack of agreement between objective and subjective measures in the evaluation of masticatory function: a preliminary study. Physiology & Behavior. 2018; 184: 220–225.
- [76] Pereira LJ, Gaviao MBD, Bonjardim LR, Castelo PM, van der Bilt A. Muscle thickness, bite force, and craniofacial dimensions in adolescents with signs and symptoms of temporomandibular dysfunction. European Journal of Orthodontics. 2007; 29: 72–78.
- [77] Rentes AM, Gavião MB, Amaral JR. Bite force determination in children with primary dentition. Journal of Oral Rehabilitation. 2002; 29: 1174– 1180.
- [78] Sakashita R, Inoue N, Kamegai T. Can oral health promotion help develop masticatory function and prevent dental caries? Community Dent Health. 2006; 23:107–115.
- [79] Sato N, Yoshiike N. Dietary patterns affect occlusal force but not masticatory behavior in children. Journal of Nutritional Science and Vitaminology. 2011; 57: 258–264.

- [80] Sonnaville WFC, Speksnijder CM, Zuithoff NPA, Verkouteren DRC, Wulffraat NW, Steenks MH, *et al.* Maximum bite force in children with juvenile idiopathic arthritis with and without clinical established temporomandibular joint involvement and in healthy children: a crosssectional study. Journal of Oral Rehabilitation. 2021; 48: 774–784.
- [81] Sonnesen L, Bakke M. Molar bite force in relation to occlusion, craniofacial dimensions, and head posture in pre-orthodontic children. European Journal of Orthodontics. 2005; 27: 58–63.
- [82] Szymańska J, Sidorowicz L. Bite force and its correlation with long face in children and youth. Folia Morphologica. 2015; 74: 513–517.
- [83] Takeshima T, Fujita Y, Maki K. Factors associated with masticatory performance and swallowing threshold according to dental formula development. Archives of Oral Biology. 2019; 99: 51–57.
- [84] Thongudomporn U, Chongsuvivatwong V, Geater A. The effect of maximum bite force on alveolar bone morphology. Orthodontics & Craniofacial Research. 2009; 12: 1–8.
- ^[85] Gudipaneni RK, Alam MK, Patil SR, Karobari MI. Measurement of the maximum occlusal bite force and its relation to the caries spectrum of first permanent molars in early permanent dentition. Journal of Clinical Pediatric Dentistry. 2020; 44: 423–428.
- ^[86] Guo R, Hama Y, Hosoda A, Kubota C, Minakuchi S. Age and sex

differences in oral functions from junior high school to young adulthood: a cross-sectional study. Journal of Oral Rehabilitation. 2021; 48: 1373– 1379.

- [87] Aishwarya N, Nagarathna C, Poovani S, Thumati P. Comparison of bite force and the influencing factors pre- and post-cementation of stainless steel crown in children using T-scan. International Journal of Clinical Pediatric Dentistry. 2021; 14: 46–50.
- [88] Prabahar T, Gupta N, Chowdhary N, Sonnahalli NK, Chowdhary R, Reddy VR. Comparative evaluation of occlusal bite force in relation to the muscle activity in the mixed dentition children of age group 9–12 years: a T-scan analysis. International Journal of Clinical Pediatric Dentistry. 2021; 14: S29–S34.

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